## An introduction to Computational Anatomy

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September 2018

#### Outline

1. Introduction to Computational Anatomy

- 2. Registration technics
- 3. Statistical analysis of the deformations
- 4. Bayesian Modelling for template estimation

#### Outline

- 1. Introduction to Computational Anatomy
  - Goals
  - Mathematical tools
  - Databases
  - Deformable Template framework
- 2. Registration technics
- 3. Statistical analysis of the deformations
- 4. Bayesian Modelling for template estimation

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- Learn the temporal evolution (growth, evolution of a disease):
   Model organ development across time
- Segmentation, prediction, help therapy:
   Build prior knowledge to simulate new anatomies, segment areas or organs in new patient, predict from the shape of an organ the evolution of a disease (through clinical variables for example)

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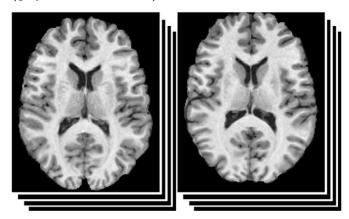
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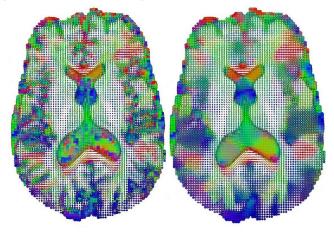
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- Etc...

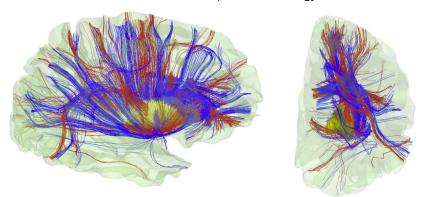
• Images (grey level, tensors, etc)



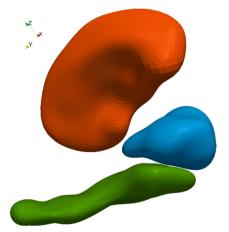
• Images (grey level, tensors, etc)



• Anatomical landmarks : anatomical points, fibres, gyri, etc



• Meshed surfaces (with point correspondence or not)



• Clinical variables related to the data (age, diagnosis, physiological parameters, etc)

Alzheimer Disease Assessment Scale— Cognitive Subscale (ADAS-cog) 11-Item	
	Score range
emory and new learning	0 - 35
Nord recall (mean number of words not recalled)	0 - 10
Orientation (one point for each incorrect response)	0 - 8
Word recognition (mean number of incorrect responses)	0 - 12
Remembering test instructions	0 - 5
nguage	0 - 25
Commands	0 - 5
Spoken language ability	0 - 5
Naming objects/fingers	0 - 5
Word-finding difficulty	0 - 5
Comprehension	0 - 5
axis	0 - 10
Constructional praxis	0 - 5
deational praxis	0 - 5
tal	0 - 70

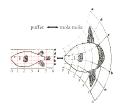
## NEED: compare these elements in a mathematical way

--- Computational anatomy is the correct setting

 $\longrightarrow$  Requires models of the data

Deformable Template Model (Genander, '80)

 Compare two observations via the quantification of the deformation from one to the other (D'Arcy Thompson, 1917)

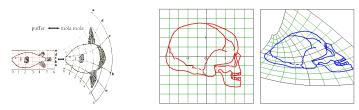






• Each element of a population is a smooth deformation of a template

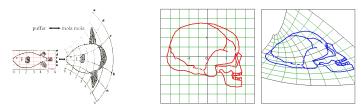
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Registration

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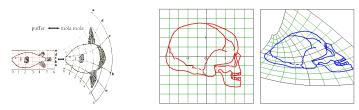


Registration

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Template estimation

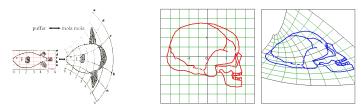
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Registration / Variance

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  - The Registration issue
  - First approach : Rigid body transformations
  - Next step : Linear Models
  - The diffeomorphic setting
- 3. Statistical analysis of the deformations
- 4. Bayesian Modelling for template estimation

### The Registration issue

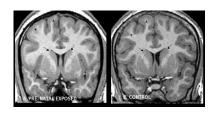




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$$I_1 \simeq \phi$$
 .  $I_O$ 

### The Registration issue





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• How to quantify the difference between the two objects?



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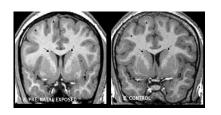
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- How to quantify the difference between the two objects?
- What kind of deformations?
- How to apply the deformation to an object?



## Object distances

Difference between two objects : (Same for the following solutions)

•  $Images: L^2$  difference of the two functions

$$||I_1 - \phi \cdot I_0||_2^2$$

Landmarks : Sum of the Euclidean distance between points :

$$\sum_{1 \le i \le N} \|x_i^1 - \phi \cdot x_i^0\|_2^2$$

 Unlabeled landmark set, meshes, fibers: Requires to embed the objects into a mathematical space where "addition, mean" and other mathematical operations are stable (see Jean)!



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- Application of the deformation to :

Images : 
$$\phi$$
 .  $I_0 = I_0 \circ \phi^{-1}$ 

 $\rightarrow$  Image support is deformed, grey levels are transported

Landmarks : 
$$\phi$$
 .  $(x_i)_{1 \le i \le N} = (\phi(x_i))_{1 \le i \le N}$ 

Others: see Jean!



# Rigid body registration

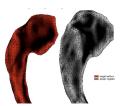


ullet Easy parametrisation of  $\phi o$  Fast computations



Restricted deformations





No affine transformation to match these objects : Need for non linear deformations

• Same difference between two objects

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- Idea: each pixel/voxel/point has its own movement

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Images : 
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Landmarks : 
$$\phi \cdot (x_i)_{1 \le i \le N} = (x_i + v(x_i))_{1 \le i \le N}$$





- ullet  $\phi$  only depends on  $v o \mathsf{Explicit}$  parametrisation
- Relevant for small deformations (where  $Id v \simeq \phi^{-1}$  is valid)



 No invertibility guaranteed : overlaps may appear → some tissue may disappear



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- ullet No invertibility guaranteed : overlaps may appear o some tissue may disappear
  - ullet Need of a diffeomorphic condition on  $\phi$



#### LDDMM framework

Let  $\Omega \in \mathbb{R}^d$  be an open set.

The framework defines

- a class of **objects**  $\mathcal{O}$  (eg : images  $I:\Omega\to\mathbb{R}$ , landmarks  $(x_i)_{1\leq i\leq N},...$ )
- a class of diffeomorphic deformations  $\phi: \Omega \to \Omega$ ,  $\mathcal{D}$
- a specific group action

• At each time step the deformation is a "small deformation" :

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- When  $\Delta t \rightarrow 0$ ,  $\phi_t$  solution of :

The flow equation : 
$$\left\{ \begin{array}{lcl} \frac{d\phi_t^v}{dt} &=& v_t \circ \phi_t^v \\ \phi_0^v &=& \mathit{Id} \end{array} \right.$$



If  $v \in V$ , with V "admissible space", then :

• Existence and uniqueness of the solution of the flow equation

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- and a **distance** between two objects  $O_0$  and  $O_1$  is computed via the group action :

$$d(O_0, O_1) = \inf_{v_t \in V, \phi_1^v(O_0) = O_1} d(Id, \phi_1^v)$$



• Existence, uniqueness and regularity (diffeomorphism) of the flow  $\phi^{\nu}$  guaranteed for  $(v_t)_{t\in[0,1]}$  such as

$$\int_0^1 \|v_t\|_V^2 dt < \infty .$$

• Then :  $d(O_0, O_1) = \inf_{v_t \in V, \phi_v^v(O_0) = O_1} \{ \int_0^1 ||v_t||_V^2 dt \}$ 

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- A distance between diffeomorphic objects: deformation cost
- A similarity term between objects: data attachment term

An energy to minimise:

$$E(v) = \underbrace{\frac{1}{2} \int_{0}^{1} \|v_{t}\|_{V}^{2} dt}_{\text{deformation cost}} + \underbrace{\frac{1}{2\sigma^{2}}}_{\text{tradeoff parameter}} \times \underbrace{|I_{1} - I_{0} \circ (\phi_{1}^{v})^{-1}|^{2}}_{\text{data attachment term}}$$

# The Hamiltonian formulation :

Let V be a RKHS which kernel  $K_V$ : (example for N landmarks)

$$ullet$$
  $\exists$   $(p_i(t))_{1\leq i\leq N}\in\mathbb{R}^n$  such that  $v_t(x)=\sum\limits_{i=1}^N K_V(x_i(t),x)p_i(t)$ 

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Then

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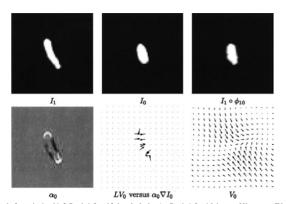
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• Hamiltonian constant on the geodesics

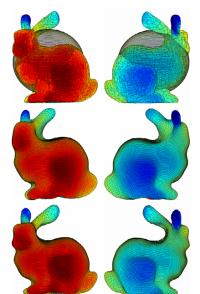


# Consequences:

- $\Rightarrow$  Parametrisation of the problem by only  $x_0$  (known) and  $p_0$
- For images : the momentum supported by the gradient of the image



# LDDMM framework : Example of deformations



#### So now?

Can do statistics on either the objects or the deformations! For example :

- Global shape analysis
- Local deformation pattern detection

#### Outline

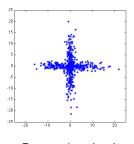
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  - Noisy ICA general model
  - Gaussian Graphical Models
- 4. Bayesian Modelling for template estimation

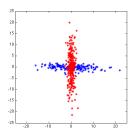
## Noisy ICA: Goal of the decomposition:



Data point cloud

- Data : one point cloud
- Goal : explain these data

### Noisy ICA: Goal of the decomposition:



Data point cloud Mixture of two Gaussian distributions

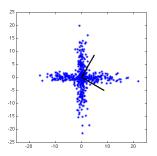
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• Goal : explain these data

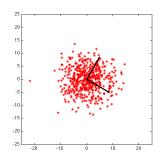
• How: Extracting the "sources" which had generated the data

## Principal Component Analysis (PCA)

- Orthogonal direction of maximum variance :
- Assumes a **Gaussian** distribution :  $X = \mu + \Sigma^{1/2} \varepsilon$ ,  $\varepsilon \sim \mathcal{N}(0, Id)$ .



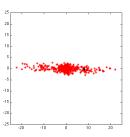
Result from PCA (black lines)



Resampling from the model

## Independent Component Analysis (ICA)

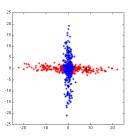
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- Finds sources which may have generated the cloud
- Accounts for non Gaussian distributions (more flexible model)



First estimated source

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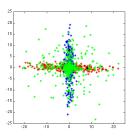
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First estimated source Second estimated source Resampling form the model

### Difference between PCA and ICA

PCA	ICA		
Maximum variance axes	Source separation		
Geometrical (orthogonal axis)	Statistical (source points in the plane)		
<b>Description</b> of data	Explanation and interpretation		
Gaussian distribution	Many other possible distributions		
only	e.g. mixtures, continuous or discrete		
	(see later)		
25 20 15 10 	25 20 15 10 3 40 115 20 25 40 10 20 : > > > > > > > > > > > > > > > > > >		

- \* Observations :  $X_1^n$  in  $(\mathbb{R}^d)^n$
- \* Source matrix : A called decomposition matrix
- st Gaussian noise :  $\sigma arepsilon_i$ 
  - \* Independent components :  $oldsymbol{eta}_i \in \mathbb{R}^p$ , p << d
  - → random vector with **independent** coordinates

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  - $\rightarrow$  random vector with **independent** coordinates  $\beta_1^n$  hidden variables.
  - Model : for all images  $X_i$ ,  $1 \le i \le n$

$$\left\{ \begin{array}{ll} \boldsymbol{\beta}_{i,j} & \sim & \nu_{\eta} \mid \eta \,, \forall \,\, 1 \leq j \leq p \\ \\ \boldsymbol{X}_{i} & \sim & \mathcal{N}(\boldsymbol{A}\boldsymbol{\beta}_{i}, \sigma^{2}\boldsymbol{Id}) \,\, \mid \boldsymbol{A}, \,\, \sigma^{2}, \,\, \boldsymbol{\beta}_{i} \,. \end{array} \right.$$

• Various choices of the distribution  $\nu_{\eta}$  on the independent components



### Various examples of distributions:

# Independent Factor analysis

$$\begin{cases} \beta_{i,j} \sim \nu_{\eta} \mid \eta, \forall \ 1 \leq j \leq p \\ X_{i} \sim \mathcal{N}(A\beta_{i}, \sigma^{2}Id) \mid A, \ \sigma^{2}, \ \beta_{i} \end{cases}$$

- For identifiability,  $\nu_{\eta}$  cannot be Gaussian
- $\nu_{\eta}$  is a mixture of K 1D Gaussian distributions  $\mathcal{N}(m_k, 1), \ k = 1, ..., K$  with weights  $(w_k)_{1 \leq k \leq K}$ .
- $\bullet \ \eta = (m_k, w_k)_{1 \leq k \leq K}$
- $\bullet \ \theta = (A, \sigma^2, (m_k, w_k)_{1 \le k \le K})$



### Various examples of distributions :

#### Continuous distributions

$$\left\{ \begin{array}{ll} \boldsymbol{\beta}_{i,j} & \sim & \boldsymbol{\nu_{\eta}} \mid \boldsymbol{\eta} \,, \forall \,\, 1 \leq j \leq p \\ \\ \boldsymbol{X_i} & \sim & \mathcal{N}(\boldsymbol{A}\boldsymbol{\beta}_i, \sigma^2 \boldsymbol{Id}) \,\, \mid \boldsymbol{A}, \,\, \sigma^2, \,\, \boldsymbol{\beta}_i \end{array} \right.$$

- $\nu_{\eta}$  is either :
  - Logistic Log(1/2),
  - Laplacian,
  - Exponentially scaled Gaussian(EG):  $\beta_i^j = s_i^j Y_i^j$  where  $\mathbf{Y} \sim \mathcal{N}(\boldsymbol{\mu}, Id)$  and  $\boldsymbol{\mu} = (\mu, \dots, \mu)$ ;  $s_i^1, \dots, s_i^p$  are independent  $\mathcal{E}xp(1)$ , also independent from  $\mathbf{Y}$  (sub-exponential tail)
- $\eta = \emptyset$  or  $\mu$
- $\theta = (A, \sigma^2)$  or  $\theta = (A, \sigma^2, \mu)$



### Various examples of distributions :

#### Discrete distributions

$$\begin{cases} \beta_{i,j} \sim \nu_{\eta} \mid \eta, \forall \ 1 \leq j \leq p \\ X_{i} \sim \mathcal{N}(A\beta_{i}, \sigma^{2}Id) \mid A, \ \sigma^{2}, \ \beta_{i} \end{cases}$$

Idea: Introduce a switch to cancel some of the decomposition vectors

→ Either binary ("on/off") or ternary (activate, inhibit, remove)

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Idea: Introduce a switch to cancel some of the decomposition vectors

- → Either binary ("on/off") or ternary (activate, inhibit, remove)
  - Bernoulli-censored Gaussian (BG) :  $\beta^j = b^j Y^j$  with  $b^j \sim \mathcal{B}(\alpha)$ , **Y** is a Gaussian vector with distribution  $\mathcal{N}(\mu, Id)$ .
  - Exponentially scaled Bernoulli-censored Gaussian (EBG): mix of EG and BG
  - Exponentially-scaled ternary distribution (ET) :  $\beta^j = s^j Y^j$ , where  $s^1, \ldots, s^p$  are i.i.d.  $\mathcal{E}xp(1)$ .  $\gamma = P(Y^j = -1) = P(Y^j = 1)$ , providing a symmetric distribution for the components of  $\mathbf{Y}$ .
  - $\theta = (A, \sigma^2, \mu, \alpha, \gamma)$

4□ > 4個 > 4 = > 4 = > = 900

#### ML Estimator:

ullet Parameters heta are estimated by maximum likelihood :

$$\hat{\theta} = \arg\max P(X; \theta)$$

•  $\beta_1^n$  unobserved random variables + Maximise a likelihood  $\rightarrow$  **EM algorithm** (Expectation - Maximisation)

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Iteration k of the algorithm :

**E Step**: Compute the posterior distribution of  $\beta_i$  given  $X_i$  and  $\theta_l$ :  $\nu_{i,k}(d\beta_i)$ 

**M Step :** Parameter update :

$$\theta_{k+1} = \arg\max_{\theta} \mathbb{E}_{\nu_{1,k}^n} \left[ \log q(\beta_1^n, X_1^n; \theta) \right].$$

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BUT: E step not computationally tractable!

#### Others:

### FastICA:

• Employs a fixed point algorithm to minimise the mutual information between the coordinates of  $A^{-1}X$ 

### Particle filtering within EM:

Approximates the posterior distribution using particle filtering

# Experiments

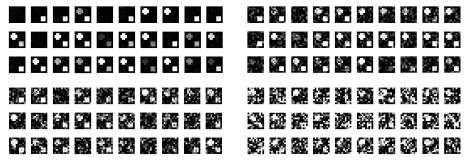


Two source images

### Experiments



Two source images



Samples of the four training sets different level of noise. From left to right and top to bottom :  $\sigma = 0.1, 0.5, 0.8, 1.5$ 

### Results of the PCA decomposition

$$\sigma = 0.1$$



$$\sigma = 0.8$$



Cumulative eigen values of the PCA decomposition

















Two first Principal Components (orthogonal images).

## Comparison: 30 images per training set

Model/Algo  $\sigma = 0.1$  $\sigma = 0.5$  $\sigma = 0.8$  $\sigma = 1.5$ FAM-EM/Log SAEM/Log SAEM/IFA EM/IFA SAEM/BG **FastICA** 

# Comparison : 50 images per training set

Model/Algo	$\sigma = 0.1$	$\sigma = 0.5$	$\sigma = 0.8$	$\sigma=1.5$
FAM-EM/Log		<b>22</b>	经团	
SAEM/Log	• •	• 4	*	
SAEM/IFA	• •	· •		
EM/IFA	٠. ٠.		5.	
SAEM/BG	• •	•	* ·	e To
FastICA	. *.			

# $Comparison: 100 \ images \ per \ training \ set$

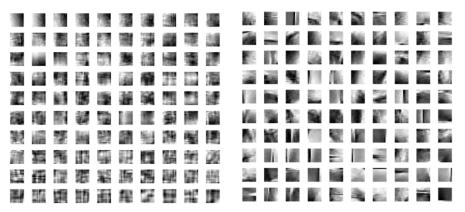
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FAM-EM/Log				资 怨
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#### Patches of faces from Caltech101 database



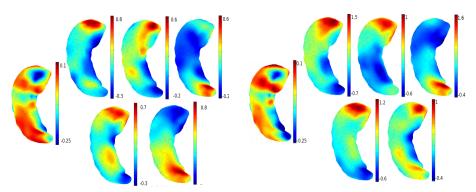
100 random images picked from the 10,000 images used as the training set. These images are patches extracted from the face images of the Caltech101 data base. Each image is a grey level image of size  $13 \times 13$ .

#### Patches of faces from Caltech101 database



100 decomposition vectors from 2 models. Left: Log-ICA. Right: BG-ICA.

### 101 hippocampus deformations (3 populations : Ctrl - Mild AD - AD)



Mean and five decomposition vectors estimated with L-ICA (left) and ET-ICA (right). Each image has its own colorbar to highlight the major patterns.

## 101 hippocampus deformations (3 populations : Ctrl - Mild AD - AD)

	Ctrl/AD		Ctrl/mild AD	
Model	L-ICA	BG-ICA	L-ICA	BG-ICA
		$0.33 \times 10^{-3}$		
Std dev.	$0.16 \times 10^{-3}$	$0.25 \times 10^{-3}$	$3.8 \times 10^{-3}$	$4.6 \ 7.6 \times 10^{-3}$

TABLE – Mean and standard deviation of the p-values for the two models with the decomposition vectors. Means and standard deviations are computed over 50 runs to separate the Controls from the AD group (left columns) and to separate the Controls from the mild AD group (right columns). PCA p-values :  $0.3 \times 10^{-3}$  and  $7.7 \times 10^{-3}$  using 95% of the cumulative variance.

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Correlations describe the global statistical dependencies between variables

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### However:

- Correlations describe the global statistical dependencies between variables
  - = both direct and indirect interactions

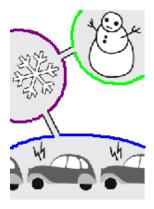
### Correlation vs Conditional Correlation



# Under Gaussian assumption

 Traffic jam intensity correlated to Number of snowmen in town due to snowstorm.

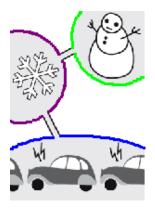
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  - ⇔ No edge between Traffic jam and Snowmen random variables

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Significant edges only appear and estimation more stable

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ullet May be too strong constrain (numerically un-invertible)  $\longrightarrow$  alleviate through :

$$X_a - X_{m_a} (X_{m_a}^T X_{m_a} + \gamma_0 Id)^{-1} X_{m_a}^T X_a.$$
 (2)



### No prior

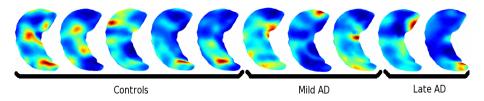
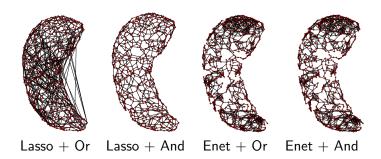


FIGURE – Examples of the training set. The colour depends on the intensity of the Jacobian of the deformation. Blue means a contraction and red dilatation. The intensity itself is not important but rather its relative value with respect to the others.

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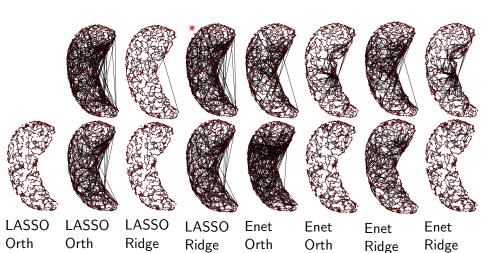


## With prior



FIGURE – Two examples of neighbourhood graphs we used. Left : 3nearest-neighbour graph. Right : neighbours have Euclidean distance below a given threshold.

Note: User's choice: you can add connection that you know.





Or

And

Or

Or

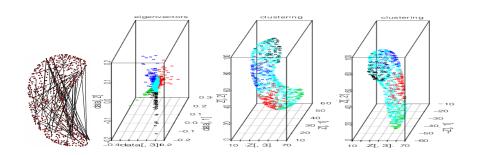
Or

And

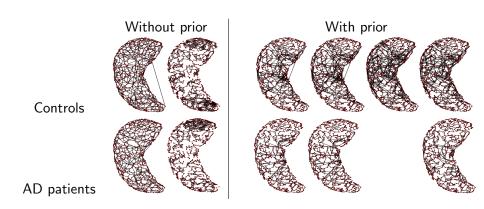
And

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## Clustering of the Shape: using spectral clustering



## Population comparison



### Outline

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- 2. Registration technics
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  - Mathematical framework for deformable models
  - Past approaches to compute a population average
  - Generative statistical models
  - Statistical estimation of the model parameters
  - Experiments
  - And next?

## What are the mean and the variability

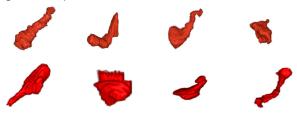
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## Other question (related)

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such that

$$y(s) = I_t(x_s - z(x_s)) + \epsilon_t(s) = z \cdot I_t(s) + \epsilon(s),$$

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  - $I_t \in V_p$  RKHS with kernel  $K_p$
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- $z \in V_g$  RKHS with kernel  $K_g$
- Given  $(g_k)_{1 \le k \le k_g} \exists (\beta^{(1)}, \beta^{(2)}) \in \mathbb{R}^{k_g} \times \mathbb{R}^{k_g}$  such that :

$$z(x) = (\mathbf{K}_{\mathbf{g}}\beta)(x) = \sum_{k=1}^{k_g} K_g(x, g_k)(\beta^{(1)}(k), \beta^{(2)}(k)).$$

Hierarchical generative model :

$$\begin{cases} \rho \sim \nu_{\rho}, \\ \theta = (\alpha_{t}, \sigma_{t}^{2}, \Gamma_{g}^{t})_{1 \leq t \leq T} \sim \otimes_{t=1}^{T} (\nu_{\rho} \otimes \nu_{g}) \end{cases}$$

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weights prior

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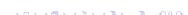
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draw images from  $\beta_1^n$  and  $\theta_{\tau}$ 

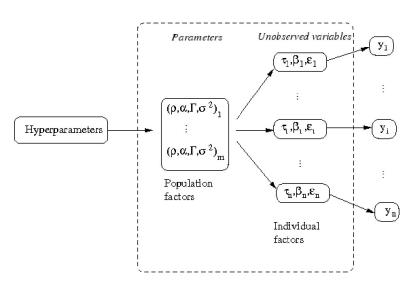


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+ weakly informative priors





#### How to learn the parameters? the MAP Estimator:

Parameters  $\theta$  are estimated by maximum posterior likelihood :

$$\hat{\theta} = \arg\max P(\theta|y)$$

where  $\theta \in \Theta = \{ (\alpha, \sigma^2, \Gamma_g) | \alpha \in \mathbb{R}^{k_p}, \ \sigma^2 > 0, \ \Gamma_g \in \mathcal{S}ym_{2k_g, *}^+(\mathbb{R}) \}.$  $\mathcal{S}ym_{2k_g, *}^+(\mathbb{R})$  is the set of positive definite symmetric matrices.

Let  $\Theta_* = \{ \theta_* \in \Theta \mid E_P(\log q(y|\theta_*)) = \sup_{\theta \in \Theta} E_P(\log q(y|\theta)) \}$  where P denotes the distribution governing the observations.

### How to do in practice?

Since  $\beta_1^n$  are unobserved variables, a natural approach to reach the MAP estimator is the **EM algorithm**.

Iteration I of the algorithm:

**E Step :** Compute the posterior law on  $\beta_i$ , i = 1, ..., n.

**M Step:** Parameter update:

$$\theta_{l+1} = \arg\max_{\theta} E\left[\log q(\theta, \beta_1^n, y_1^n)|y_1^n, \theta_l\right].$$

BUT: the E step is not tractable!

## E step: First solution proposed:

## Fast approximation with modes:

E Step :

 $\nu_{i,k}(d\beta_i) \simeq \delta_{\beta_i^*}$ ,  $\forall i = 1, ..., n$ .  $\beta_i^*$  maximise the conditional distribution on  $\beta$  with the current parameters :

$$\beta_i^* = \arg\max_{\beta} \log q(\beta|X_i; \theta_k)$$

M Step: Parameter update: uses the "completed observations"

$$\theta_{k+1} = \arg\max_{\theta} \log q((\beta^*)_1^n, X_1^n; \theta).$$

## Details of the maximisation step :

#### Geometry:

$$\theta_{g,l+1} = \Gamma_{g,l+1} = \frac{1}{n+a_g} (n[\beta \beta^t]_l + a_g \Sigma_g).$$

where

$$[\beta\beta^t]_I = \frac{1}{n} \sum_{i=1}^n \int \beta\beta^t \nu_{I,i}(\beta) d\beta,$$

is the empirical covariance matrix with respect to the posterior density function.

 $\rightarrow$  Importance of the prior!

### Details of the maximisation step:

## Photometry:

$$\begin{cases} \alpha & = \left( n \left[ \left( K_{p}^{\beta} \right)^{t} K_{p}^{\beta} \right]_{I} + \sigma^{2} \Sigma_{p}^{-1} \right)^{-1} \left( n \left[ \left( K_{p}^{\beta} \right)^{t} Y \right]_{I} + \sigma^{2} \Sigma_{p}^{-1} \mu_{p} \right) \\ \sigma^{2} & = \frac{1}{n |\Lambda| + a_{p}} \left( n \left[ Y^{t} Y \right]_{I} + \alpha^{t} \left[ \left( K_{p}^{\beta} \right)^{t} K_{p}^{\beta} \right]_{I} \alpha \\ & -2 \alpha^{t} \left[ \left( K_{p}^{\beta} \right)^{t} Y \right]_{I} \right) + a_{p} \sigma_{0}^{2} \right) . \end{cases}$$

# E step: First solution proposed: Fast approximation with modes:

•  $\nu_{i,i}^*(d\beta_i) = \delta_{\beta_i^*}$ ,  $\beta_i^*$  maximise the conditional distribution on  $\beta$  for each component:

$$\begin{split} \beta_{i,\tau}^* &= \arg\max_{\beta} \log q(\beta | \alpha_{l,\tau}, \sigma_{l,\tau}, \Gamma_{g,l,\tau}, y_i, \tau) = \\ &\arg\min_{\beta} \left\{ \frac{1}{2} \beta^t (\Gamma_{g,l,\tau})^{-1} \beta + \frac{1}{2\sigma_{l,\tau}^2} |y_i - \textit{K}_p^\beta \alpha_{l,\tau}|^2 \right\} \;, \end{split}$$

- Maximise the conditional distribution on  $\tau$  given the  $\beta_i^*$ .
- Pick  $\beta_{i,\tau^*}^*$ .

# Advantages and drawbacks:



- Computation of  $\beta_i^*$ : standard gradient descent.
- Reduce the EM algorithm to an iterative maximisation of the joint density.



• Highly sensitive to noise (see experiments)



## Our solution: MCMC- Stochastic Approximation EM algorithm:

Iteration  $k \to k+1$  of the algorithm :

- Simulation step :  $\beta^{k+1} \sim \Pi_{\theta_l}(\beta^l, \cdot)$  where  $\Pi_{\theta_k}(\beta^k, \cdot)$  is a transition probability of a convergent Markov Chain having the posterior distribution as stationary distribution,
- Stochastic approximation :

$$Q_{k+1}(\theta) = Q_k(\theta) + \Delta_k[\log q(\boldsymbol{X}, \boldsymbol{\beta}^{k+1}; \theta) - Q_k(\theta)]$$

where  $(\Delta_k)$  is a decreasing sequence of positive step-sizes.

• Maximisation step :  $\theta_{k+1} = \arg \max Q_{k+1}(\theta)$ 

[\*]  $\Pi_{\theta_{k}}(\beta^{l},\cdot)$  given by different samplers.



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# Our solution: MCMC- Stochastic Approximation EM algorithm: (2):

#### But!

All our models belong to the Exponential family,

$$q(\mathbf{X}, \boldsymbol{\beta}^{k+1}; \boldsymbol{\theta}) = \exp \left\{ -\psi(\boldsymbol{\theta}) + \langle S(\mathbf{X}, \boldsymbol{\beta}), \phi(\boldsymbol{\theta}) \rangle \right\}$$

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Very simple algorithm!



With these models and algorithms we have proved some important asymptotic results :

### Conditions:

- Smoothness of the model (classic conditions for convergence of stochastic approximation and EM)
- In case of AMALA sampler : condition for its geometric ergodicity

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- Convergence of  $(s_k)$  a.s. towards critical point of mean field of the problem
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- Central limit theorem for  $(\theta_k)$



### Training sets

FIGURE – Left : Training set (inverse video). Right : Noisy training set (inverse video).

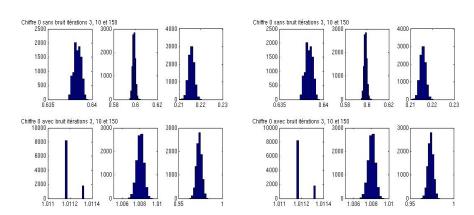
# MCMC-SAEM algorithm :

Algorithm/ Noise level	FAM-EM	H.GSAEM	AMALA-SAEM
No Noise	0   2 3 4	01234	01234
	5 6 7 8 9	56789	56789
Noisy	01234	01234	01234
of Variance 1	56789	56789	56789

FIGURE – Estimated templates using different algorithms and two level of noise. The training set includes 20 images per digit.

# Estimated geometric variability

000000000000000000000000000000000000000	000000000000000000000000000000000000000
11111111111111111111111111111111111111	22 22 22 22 22 22 22 22 22 22 22 22 22
333333333333333333333333333333333333333	33333333333333333333333333333333333333
44444444444444444444444444444444444444	44444444444444444444444444444444444444
66666666666666666666666666666666666666	66666666666666666666666666666666666666
77777777777777777777777	7777777777777777777777777 888888888888
\$\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\	4949494494494494494 949449494999999



 ${f Figure}$  – Evolution of the estimation of the noise variance along the SAEM iterations. Test of convergence towards the Gaussian distribution of the estimated parameters.

### Classification rates:

Error rate	"EM-Mode" est.	SAEM-MCMC est.
Mode classifier	40.71	22.52
MCMC classifier	-	17.07

TABLE - Error rate with respect to the estimation and classification methods.

### 3D dendrite spines:





FIGURE – Estimated template with the one component model : Left : 3D representation of the grey level volume. Right : 3D representation of the thresholded volume.

# 3D dendrite spines:



FIGURE – Estimated templates of the two components with the 30 image training set : 3D representation after thresholding.

# 3D dendrite spines:

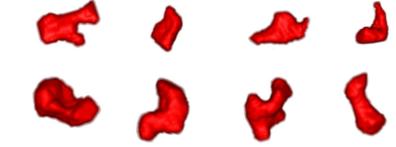


FIGURE - 3D view of eight synthetic data. The estimated template shown in Figure 7 is randomly deformation with respect to the estimated covariance matrix. The results are then thresholded in order to get a binary volume.

One step further?

Models of longitudinal data!